Contents lists available at SciVerse ScienceDirect

Science of the Total Environment

journal homepage: www.elsevier.com/locate/scitotenv

Pathogen reduction requirements for direct potable reuse in Antarctica: Evaluating human health risks in small communities



S. Fiona Barker^{a,b,*}, Michael Packer^c, Peter J. Scales^d, Stephen Gray^e, Ian Snape^c, Andrew J. Hamilton^f

^a Department of Resource Management and Geography, The University of Melbourne, Parkville, VIC, 3010 Australia

^b Department of Primary Industries Victoria, Parkville, VIC, 3052 Australia

^c Department of Sustainability, Environment, Water, Population and Communities, Australian Antarctic Division, 203 Channel Highway, Kingston, Tasmania 7050 Australia

^d Particulate Fluids Processing Centre, Department of Chemical and Biomolecular Engineering, The University of Melbourne, Parkville, VIC 3010 Australia

^e Institute for Sustainability Innovation, Victoria University, PO Box 14428, Melbourne, VIC, 8001 Australia

^f Department of Agriculture and Food Systems, The University of Melbourne, Dookie College, VIC, 3647 Australia

HIGHLIGHTS

GRAPHICAL ABSTRACT

- Direct potable reuse (DPR) projects should consider population size.
- Small community pathogen load in outbreak sewage is higher (p<0.001) than municipal.
- LRVs for municipal sewage: 6.9 (norovirus), 8.0 (giardia), 7.4 (Campylobacter).
- LRVs for small community: 12.1 (norovirus), 10.4 (giardia), 12.3 (Campylobacter).
- Additional treatment barriers required for small community DPR to meet 10–6 DALYs.



ARTICLE INFO

Article history: Received 15 April 2013 Received in revised form 20 May 2013 Accepted 20 May 2013 Available online 14 June 2013

Editor: D. Barcelo

Keywords: Campylobacter Drinking water Giardia

ABSTRACT

Small, remote communities often have limited access to energy and water. Direct potable reuse of treated wastewater has recently gained attention as a potential solution for water-stressed regions, but requires further evaluation specific to small communities. The required pathogen reduction needed for safe implementation of direct potable reuse of treated sewage is an important consideration but these are typically quantified for larger communities and cities. A quantitative microbial risk assessment (QMRA) was conducted, using norovirus, giardia and *Campylobacter* as reference pathogens, to determine the level of treatment required to meet the tolerable annual disease burden of 10^{-6} DALYs per person per year, using Davis Station in Antarctica as an example of a small remote community. Two scenarios were compared: published municipal sewage pathogen loads and estimated pathogen loads during a gastroenteritis outbreak. For the municipal sewage scenario, estimated required \log_{10} reductions were 6.9, 8.0 and 7.4 for norovirus, giardia and *Campylobacter* respectively, while for the outbreak scenario the values were 12.1, 10.4 and 12.3 (95th percentiles). Pathogen concentrations are higher under outbreak conditions as a function of the relatively

Abbreviations: DALYs, disability adjusted life years; DPR, direct potable reuse; IPR, indirect potable reuse; LRV, log10 reduction values; QMRA, quantitative microbial risk assessment.

* Corresponding author at: Department of Resource Management and Geography, The University of Melbourne, Parkville, VIC 3010, Australia. Tel.: +61 3 8341 2413.

E-mail addresses: fionabr@unimelb.edu.au (S.F. Barker), Michael.Packer@aad.gov.au (M. Packer), peterjs@unimelb.edu.au (P.J. Scales), Stephen.Gray@vu.edu.au (S. Gray), Ian.Snape@aad.gov.au (I. Snape), andrewjh@unimelb.edu.au (A.J. Hamilton).

0048-9697/\$ – see front matter ${\odot}$ 2013 Elsevier B.V. All rights reserved. http://dx.doi.org/10.1016/j.scitotenv.2013.05.059





Norovirus Quantitative Microbial Risk Assessment (QMRA) Sewage greater degree of contact between community members in a small population, compared with interactions in a large city, resulting in a higher proportion of the population being at risk of infection and illness. While the estimates of outbreak conditions may overestimate sewage concentration to some degree, the results suggest that additional treatment barriers would be required to achieve regulatory compliance for safe drinking water in small communities.

© 2013 Elsevier B.V. All rights reserved.

1. Introduction

Small remote communities sometimes struggle to adequately meet basic services such as power and water. In Australia, for example, there are many small remote communities. This is exemplified by the many remote indigenous communities, with nearly 13% of people living in the 838 communities with a population of less than 50 people and a significant number in communities with between 50 and 199 residents (ABS, 2008). More than half of the people living in remote indigenous communities rely on bore water as their main water source, 62% rely on community generators for electricity, only 30% are connected to a town sewerage system while 28% and 3.2% use septic tanks or pit toilets, respectively and high proportions of people experience interruptions in supply of services (ABS, 2008). In some of these communities, where water scarcity is an issue of concern, alternative sources of water may be needed. While recent droughts in Australia were accompanied by a drastic rise in the domestic use of grey water (ABS, 2007a, 2010a, 2010b), alternative sources of potable water have received less attention.

Indirect potable reuse schemes for the recycling of wastewater (IPR is the discharge of treated water into a receiving body prior to extraction and re-treatment for potable use) can be found in many countries; however, direct potable reuse (DPR is reuse without environmental mixing) is rare. There are currently only three DPR schemes in the world: Windhoek in Namibia (Lahnsteiner and Lempert, 2007), Cloudcroft in New Mexico and Big Springs in Texas (Tchobanoglous et al., 2011). While the more immediate driver of DPR is extreme water scarcity, various other factors also favor DPR systems, including whole-of-system life-cycle costs, reliability of water supply and quality and the exhaustion of economically feasible non-potable reuse options (Leverenz et al., 2011). An important consideration for system design and operation is the impact of population size on disease outbreaks, sewage quality and ultimately the required level of treatment. A greater understanding of these impacts is needed before the technology is implemented broadly.

Quantitative microbial risk assessment (QMRA) is a useful tool to assess pathogen reduction requirements for wastewater recycling and has been used to inform the regulatory environment relevant to wastewater schemes for non-potable reuse, IPR and DPR scenarios (NRMMC et al., 2006b; NRMMC et al., 2008; NRMMC et al., 2009; WHO, 2006). Reuse guidelines are usually based on water quality characteristics of municipal sewage from large cities and, using a tolerable annual disease burden of $\leq 10^{-6}$ disability adjusted life years (DALYs) per person per year, QMRA has been used to inform guidelines where recommended pathogen log₁₀ reduction values (LRV) are presented (NRMMC et al., 2008). Municipal sewage is typically of consistent or relatively stable quality, as a function of the dilution effect from a large population base (NRMMC et al., 2008), although differences between peak and non-peak seasons may be detectable; for example norovirus concentrations in sewage may be up to 1 or 2 logs units higher during peak season (Katayama et al., 2008; Nordgren et al., 2009; Victoria et al., 2010). Localized disease outbreaks and changes in population size may significantly alter sewage microbial quality from a small population, potentially affecting treatment requirements.

The objective of this study was to determine the required LRVs for DPR in small communities as this has not been specifically considered in reuse guidelines. While any of a number of small remote communities could have been chosen as a representative population for the model, Davis Station, the largest of three permanent Australian research stations in Antarctica, was selected for this exercise as there is current interest in DPR. The Australian Antarctic Division is undertaking a project to reduce the environmental impact of sewage treatment and disposal at Davis Station. As part of this project, research is being conducted into the potential implementation of DPR which, in addition to providing a reliable potable water supply, could provide considerable energy savings as compared with the existing water system. While Davis Station may not be a typical small community, only minor modifications (volume of drinking water and days of exposure) would be required to adequately reflect other populations. Regardless, the results of this assessment were considered generalizable to a range of other small communities, of which there are many in Australia and around the world.

2. Methods

The focus of this model was human health risks from waterborne pathogens, in particular diarrheal diseases, from ingestion of treated drinking water. Two complementary approaches were employed to estimate sewage pathogen concentrations: published values from municipal sewage treatment plants and estimated gastroenteritis outbreak conditions. Further detail is provided in supplementary materials.

2.1. QMRA

The QMRA method was used to determine required LRVs for direct potable reuse of wastewater starting from a health target—a tolerable annual burden of disease (*DB*) of $\leq 10^{-6}$ DALYs person⁻¹ year⁻¹—that has been widely adopted for both drinking water and non-potable reuse (NRMMC et al., 2006b; WHO, 2006; WHO, 2011). All model input parameters are listed in Table 1. Using the annual burden of disease calculation

$$DB = P_{\rm ill}BS_{\rm f},\tag{1}$$

the tolerable annual probability of illness (P_{ill}) was determined, where *B* is the disease burden (DALYs per case of illness) and S_f is the proportion of the population susceptible to the disease.

While country-specific estimates of disease burden (*B*) are preferred, they are often non-existent. In this model, published values from a range of countries were used. For norovirus, a Uniform distribution (Cressey and Lake, 2009; Haagsma et al., 2008; Kemmeren et al., 2006; Lake et al., 2010; Masago et al., 2006) was used to represent the range of available values and similarly using Dutch data for giardia (Havelaar, 2012; Vijgen et al., 2007) and *Campylobacter* (Havelaar, 2012; Havelaar and Melse, 2003).

Disease susceptibility (S_f) is used to exclude the proportion of the population shown to be resistant to infection. There is evidence of resistance to norovirus infection (Johnson et al., 1990; Lindesmith et al., 2003; Teunis et al., 2008) related to both histo-blood group antigens and secretor status (Le Pendu, 2006) although it has been suggested

Table 1 Model input parameters.

ParameterUnitsDistribution or point estimates ^a , [mean ^b]References and jutility	ustification
Disease burden (B) DALYs case of illness $^{-1}$	
Norovirus Uniform $(3.71 \times 10^{-4}, 6.23 \times 10^{-3}), [3.30 \times 10^{-3}]$ (Cressey and Lake	ke, 2009; Haagsma et al., 2008;
Ciardia H_{2}^{2} (2.10 + 10 ⁻³) (2.20 + 10 ⁻³) (1.20 + 10 ⁻³)	I., 2006; Lake et al., 2010; Masago et al., 2006)
Gamulabacter Uniform $(2.10 \times 10^{-2}, 12.8 \times 10^{-2})$ (Havelagr 2012)	Havelaar and Melse 2003)
Susceptibility fraction $(S_{\rm f})$ proportion	nuvenuir und meise, 2005)
Norovirus Uniform (0.8, 1.0), [0.9] (Atmar, 2010; De Soller et al., 2010	enborough and Downing, 1968; 0; Thorven et al., 2005)
Giardia, Campylobacter 1	
Exposure events (n) days year ⁻¹ Uniform (62, 121), [91.5]Total number of c ≥ 30 (AAD, 2011)	days for months with population) between 2005 and 2010
Dose-response models	
Norovirus (a + b moculum) rui beta-roisson: $\alpha_{NV} = 0.04$, $\beta_{NV} = 0.003$, $\beta_{NV} = 0.003$, $\beta_{NV} = 0.003$	08)
Giardia Exponential: r_G = Triangular(0.0044, 0.0566, 0.0199), [0.027] (Teunis et al., 199	196); min/max are 95th confidence intervals
CampylobacterFull Beta-Poisson: $\alpha = 0.024, \beta = 0.011,$ (Teunis et al., 200)	05)
$\eta_{\rm C} = 3.63 \times 10^{-9}, r_{\rm C} = 2.44 \times 10^8$	
Giardia infection:illness proportion Uniform (0.24, 0.93), [0.58] (Birkhead and Vo	ogt, 1989; Hoque et al., 2002;
(inf):iii) Logrormal (3, 1) - truncated at 2 and 6: $u = 1.05$ & $= 0.32$ (Hunter et al. 20)	0; 1 factor et al., 2010
Schijven et al. 20	011; USEPA, 2004; USEPA, 2006)
Sewage concentration –	
municipal sewage (c _{sewage})	
Norovirus PCR units L^{-1} Mixture (A, B), [3.12 × 10 ^o]; 11.1% recovery et al. In the second se	efficiency (Katayama et al., 2008) applied
$A = Lognormal(2.19 \times 10^6 \text{ G27} \times 10^6), \mu = 14.2, 0 = 0.94, \text{ to A & 6 (Addydi)}$ B = Lognormal(2.66 \times 10^6 \text{ G27} \times 10^6) = 1.11	illa et al., 2008) (Harallioto et al., 2008)
Giardia $cysts L^{-1}$ Mixture (G1, G2, G3), [2,51 × 10 ³]; (Van Den Akker et	et al., 2011)
G1 = 10^Normal(2.90, 0.56), recovery included	d in values (32-47%)
$G2 = 10^{Normal}(2.94, 0.77),$	
$G3 = 10^{\text{NDMMC}} t_2^{-1} $	2006-2)
Cumpyroducter Cu L Expension (P) # neople Discrete distribution $(min = 51 max = 106)$ [72] Daily station non	2000a) mulation in months with population >30 .
data from 2005-2	2011, $n = 601$ (AAD, 2011)
Secondary attack rate (<i>A</i> _r) proportion	
Norovirus Uniform (0.14, 0.22), [0.18] (Alfano-Sobsey et	et al., 2012; Baron et al., 1982; Götz et al., 2002;
Johansson et al., 2 Giardia	2002; ter Waarbeek et al., 2010)
Gialula Olinoini (0.17, 0.16), [0.175] (Katz et al., 2006) Camulabacter Ulinform (0.015) [0.075] (Fixas 1996: No	orkrans and Svedhem 1982. Porter and Reid 1980)
Peak shedding rate	
Norovirus (S_{NV}) copies g-feces ⁻¹ Uniform $(2.9 \times 10^{10}, 1.6 \times 10^{12}), [8.2 \times 10^{11}]$ (Atmar et al., 200	08; Chan et al., 2006; Lee et al., 2007)
Giardia (S _G) cysts person ⁻¹ day ⁻¹ Uniform ($6.42 \times 10^8, 7.05 \times 10^8$), [6.73×10^8] (Tsuchiya, 1931)	
Campylobacter (Sc)ctu g-feces ⁻¹ Uniform (10 ⁴ , 10 ⁵), [5 × 10 ⁶](Feachem et al., 1Daily disrebeal feest weight (E) α feese person ⁻¹ Uniform (200, 750) [475](Pers 2000)	1983; Lin et al., 2008)
Daily water use (W) L person ⁻¹ dav ⁻¹ Uniform (90, 174), [132] Davis Station bether	tween 2010 and 2011 (AAD, 2011; AAD, 2012)

^a Distributions: Lognormal(mean, sd), values from 1,000,000 iterations, population parameters μ and δ calculated as follows: $\mu = \ln(\overline{x}) - 0.5\ln(1 + s^2/\overline{x}^2)$, $\delta = [\ln(1 + (s^2/\overline{x}^2))]^{1/2}$, where \overline{x} is the sample mean and s^2 the sample standard deviation; mixture is a set of random values drawn from each distribution with equal weighting; normal (mean, sd); triangular (min, max, mode/most likely); uniform (min, max). ^b Mean of 1,000,000 iterations (for information purposes only).

that, due to the variation between norovirus genotypes, every person may be genetically susceptible to at least one norovirus genotype (Atmar, 2010). Since susceptibility to norovirus is uncertain, S_f was represented by a Uniform distribution accounting for a range from secretor-positive individuals (0.8; Denborough and Downing, 1968; Thorven et al., 2005) through to all individuals (1.0). Despite many years of research, there remain many questions about the mechanisms of pathogenicity, host responses to infection and immunity to giardia infections (Roxström-Lindquist et al., 2006); therefore, in this work, all individuals were assumed susceptible ($S_f = 1$). No information on susceptibility to *Campylobacter* was found so the same assumption was made.

To estimate the tolerable daily probability of illness (p_{ill}), the original equation for annual probability of illness (WHO, 2006) was used such that

$$P_{\rm ill} = 1 - (1 - p_{\rm ill})^n, \tag{2}$$

for *n* exposure events (days year⁻¹). In the model, the summer period (months where population > 30) was assumed to be the period of exposure (due to the movement of people to and from the station) and was represented by a Uniform distribution determined from Davis Station records between 2005 and 2010. The tolerable daily probability of infection (p_{inf}) was determined using published dose–response models for norovirus (Teunis et al., 2008), giardia (Teunis et al., 1996) and campylobacter (Teunis et al., 2005). Full details of dose–response models and determination of tolerable dose are provided in supplementary materials.

The tolerable pathogen concentration in treated drinking water ($c_{tolerable}$; organisms L^{-1}) was estimated from the exposure model,

$$\lambda = c_{tolerable} V, \tag{3}$$

using the estimated tolerable dose (λ) and the daily per capita water consumption (*V*; L person⁻¹ day⁻¹). Per capita water consumption at Davis Station is much higher than that of the general population (typically assumed to be 2 L day⁻¹) as humidity is very low in Antarctica. Some community members have indicated they drink much more than the recommended 4 L, with consumption of up to 6 L per day considered quite reasonable. Variability in drinking water consumption was represented using a lognormal distribution (Åstrom et al., 2007; Pintar et al., 2012; Schijven et al., 2011) with a mean daily drinking water consumption of 3 L. In studies with mean daily drinking water consumption greater than 1 L (Table S.1), standard deviations ranged from 0.8 to 1.2; therefore, the middle value (1.0) was chosen to represent variation and the distribution was truncated at the likely minimum and maximum values (2 and 6 L).

Finally, the required log_{10} reduction value (LRV) in sewage, necessary to meet tolerable drinking water quality, was calculated as

$$LRV = log_{10}(c) - log_{10}(c_{tolerable}),$$
(4)

where the pathogen concentrations in sewage (c) were estimated using two different methods: 1) published values of pathogen concentrations in municipal wastewater and 2) estimates of sewage pathogen concentrations during a gastroenteritis outbreak at Davis Station. There was no available information on concentrations of pathogens or indicator organisms in raw sewage at Davis Station.

Norovirus, giardia and *Campylobacter* concentrations in municipal wastewater (c_{sewage} ; # L⁻¹) were assumed to follow a Lognormal distribution, with values drawn from published literature (refer to supplementary materials). An estimate of outbreak conditions at Davis Station was developed, with an outbreak defined as the arrival of one

infected person. Outbreak sewage pathogen concentrations (c_0 ; # L⁻¹) were estimated using the following equations

$$c_o = \frac{(1 + PA_r)S}{WP},\tag{5}$$

$$S = S_{\rm NV}F \text{ or } S = S_{\rm C}F, \tag{6}$$

where *P* is the population on a given summer day, A_r is the secondary attack rate (proportion), *S* is the peak daily pathogen shedding rate (person⁻¹ day⁻¹), *W* is the per capita water use (L person⁻¹ day⁻¹), *S*_{NV} and *S*_C are the norovirus and *Campylobacter* shedding concentrations (# g feces⁻¹) and *F* is the daily diarrheal excretion rate (g feces person⁻¹ day⁻¹).

To represent the summer population (*P*), months were selected where the minimum number of people on station was >30, and daily population values (n = 601) were used as a discrete distribution, using data from 2005 to 2011. Daily per capita water use (*W*) was determined from monthly average population and monthly total station water use during summer months (2010–2011; AAD, 2012), with the variation represented by a Uniform distribution.

The secondary attack rate (A_r) is the proportion of people who, after contact with the original infected person, become ill (typically measured as the number of symptomatic cases). A_r was used to estimate the maximum number of people who might be ill at one time (post-arrival of the one infected person), making the unrealistic (highly conservative) assumption that all infections occurred instantaneously (rather than over a period of days or weeks). Uniform distributions were used to represent the range of published values for secondary attack rate. Various studies have reported secondary norovirus attack rates between 0.14 and 0.22 over periods of up to 14 days after the first reported case (Alfano-Sobsey et al., 2012; Baron et al., 1982; Götz et al., 2002; Johansson et al., 2002; ter Waarbeek et al., 2010). Two studies reported very similar secondary attack rates for giardia (Katz et al., 2006; Pickering et al., 1981) while a wide range (0 to 0.15) was reported for Campylobacter (Evans, 1996; Norkrans and Svedhem, 1982; Porter and Reid, 1980).

Shedding rates (S) were also represented by Uniform distributions. The only known study of giardia shedding rates (S_G; cysts per $son^{-1} day^{-1}$) was conducted with two infected individuals over a period of 7 weeks (Tsuchiya, 1931) and the maximum shedding rate from each participant was used to define the range of peak daily shedding rates. Lin et al. (2008) reported viable Campylobacter counts in feces (CFU g^{-1}) from 10 samples while Feachem et al. (1983) reported counts as high as 10^9 per g feces (minimum and maximum values used to define the distribution). Three studies (Atmar et al., 2008; Chan et al., 2006; Lee et al., 2007) reported a range of norovirus shedding concentrations (S_{NV} ; copies g-feces⁻¹) and the maximum value from each of the four sets of data was used to define the distribution. For both norovirus and Campylobacter, a uniform distribution (# g-feces⁻¹) was converted to shedding rate using an estimate of daily diarrheal fecal weight (F; g person⁻ day^{-1}). Individuals suffering from diarrhea are typically defined as having a daily stool weight in excess of 200 g and a recent study reported mean stool weights of 750 g in persons with diarrhea (Rao, 2006); a uniform distribution was used to represent fecal weights for ill individuals, making the assumption that all infected individuals have diarrhea (secondary attack rate counts only people who are symptomatic).

2.2. Population size

The premise of this model is that small communities need to be considered differently to large cities, with the assumption that outbreak conditions will be significantly different to those in a large city as a function of the relatively greater degree of contact between community members in a small population and the greater level of dilution in a municipal sewage treatment plant due to the large population served (NRMMC et al., 2008). The estimate of municipal sewage concentrations reflects "average" conditions in a large city while outbreak sewage concentration was estimated assuming that the community at Davis Station operates in a similar fashion to the confined populations assessed to determine secondary attack rates (assuming a high degree of contact between all community members). The difference in sewage concentrations during outbreaks in small or large communities is a function of the proportion of the population infected. To evaluate the impact of population size, a method was developed to estimate the likely sewage concentration, and therefore required log₁₀ reduction, during a norovirus outbreak in a large city. Norovirus was selected as the reference pathogen as the required epidemiological data were available.

In Australia, there are 0.92 cases of gastroenteritis per person per year (Hall et al., 2006) of which 10.7% are caused by norovirus (Sinclair et al., 2005). If all norovirus infections occurred simultaneously (which is highly improbable), then 9.8% of the population would be infected (~0.098 cases of novovirus infection per person per year). A more realistic scenario can be developed using the results of a Melbourne study of 600 households that reported a maximum of 2.5% of households with at least one case of norovirus per month (Sinclair pers. comm.; Sinclair et al., 2005), assuming that monthly incidence rates equate to outbreaks. Assuming four people per household, 1.4 people would be infected per household event (average value reported by Sinclair et al., 2005), and using the current Melbourne population of 4,137,432 (ABS, 2012), an estimated 36,203 people would be infected during an outbreak, or ~0.88% of the population. Applying this monthly infection rate across a whole year, there would be 0.105 cases of norovirus per person per year which is consistent with the estimated value above (0.098) and therefore a norovirus outbreak in Melbourne was conservatively assumed to infect 1% of the population. The following scenarios were compared to evaluate the magnitude of the effect of population size: municipal sewage ("average" city conditions), outbreak conditions in a large city (population >1 million) and outbreak conditions at Davis Station.

2.3. Method comparisons

The model presented herein uses a different approach to that taken by regulatory bodies. For example, a stochastic approach was used here to account for variability and uncertainty in the model while the Australian Guidelines for Water Recycling – Augmentation of Drinking Water Supplies (NRMMC et al., 2008) use a deterministic approach, conceding that stochastic analyses may provide a better understanding of uncertainty and variability where sufficient data is available. In our model, norovirus was chosen as the reference pathogen for viruses, giardia for protozoans and *Campylobacter* for bacteria, while the Guidelines use adenovirus measurements with the rotavirus dose–response model for viruses, cryptosporidium for protozoans and *Campylobacter* for bacteria. In addition, daily per capita drinking water consumption was much higher to reflect conditions at Davis Station. The differences in methods between the model used herein and that described in the Guidelines are outlined in Table S.2. The Guideline method and input parameters were used and then individual parameters were changed sequentially (detailed in Table 2, Table S.4 and Table S.5) to evaluate the impact of each change on the model output (required LRVs).

2.4. Sensitivity analysis

A sensitivity analysis, using Spearman rank order correlation coefficients, was conducted using values from the first 1000 random draws of each input distribution to identify those input parameters that had the greatest influence on the uncertainty of the model output. Input distributions were assessed to ensure there was no correlation between unrelated variables and then relevant input parameters were tested against the final model output (*LRV*). To further evaluate the impact of variation of input parameters on the magnitude of required *LRVs*, the model was run with key inputs set at discrete percentile values (5th, 50th and 95th), with no other alteration to the model; median required LRVs were reported.

2.5. Model structure and implementation

For all input parameters, a set of random values (n = 1,000,000) was drawn from the distribution and used for all model calculations. For all model outputs, the median and 90% confidence intervals were reported. Confidence intervals were estimated using the percentile method (Buckland, 1984) and values are reported as follows: 50th [5th, 95th]; single values are 95th percentile values unless otherwise indicated. Statistical differences were determined from the first 10,000 random draws from each output distribution using analysis of variance (ANOVA) and comparison of means using Tukey's HSD (Honestly Significant Difference) test. Differences were performed in 'R' version 2.12.2 (The R Foundation for Statistical Computing, 2011) and some distribution fitting was conducted in @Risk (version 5.7).

3. Results

Estimates of norovirus, giardia and *Campylobacter* concentrations in municipal sewage $(1 \times 10^7, 9 \times 10^3 \text{ and } 7.2 \times 10^3 \text{ # L}^{-1}$, respectively) were significantly lower (p < 0.0001) than those determined for Davis Station outbreak conditions, 1.4×10^{12} , 1.4×10^6 and 4.9×10^8 (Fig. 1), which had a direct effect on the required LRVs.

Table 2

Estimated required enteric virus log₁₀ reduction values (LRVs) for stepwise methodological changes from the Guideline method (NRMMC et al., 2008) to a deterministic approximation of the model using municipal sewage concentrations.

Step	LRV	Model input parameters ^a									
		V	С	В	S _f	inf:ill	d–r	n			
1.	9.4	2	8000	$1.3 \times 10^{-2} (RV)$	0.06 (RV)	0.88	RV ^b	365			
2.	12.5	2	1.02×10^{7} (95th NV)	1.3×10^{-2} (RV)	0.06 (RV)	0.88	RV ^b	365			
3.	9.8	4.8 (95th AAD)	8000	1.3×10^{-2} (RV)	0.06 (RV)	0.88	RV ^b	365			
4.	7.2	2	1.02×10^7 (95th NV)	5.94×10^3 (95th NV)	0.99 (95th NV)	NV	NV ^c	365			
5.	7.6	4.8 (95th AAD)	1.02×10^7 (95th NV)	5.94×10^3 (95th NV)	0.99 (95th NV)	NV	NV ^c	365			
6.	6.9	2	1.02×10^7 (95th NV)	5.94×10^3 (95th NV)	0.99 (95th NV)	NV	NV ^c	118 (95th AAD)			
7.	7.3	4.8 (95th AAD)	$1.02\times10^7~(95th~\text{NV})$	5.94×10^3 (95th NV)	0.99 (95th NV)	NV	NV ^c	118 (95th AAD)			

^a Model input parameters: V = daily water consumption (L person⁻¹), c = sewage pathogen concentration ($\# L^{-1}$), B = disease burden (DALYs case⁻¹), $S_f =$ susceptibility fraction, *inf:ill* = ratio of infection to illness, d-r = dose-response model, n = days of exposure per year. 95th refers to 95th percentile of the input distribution. AAD = Davis Station data. NV = norovirus, RV = rotavirus.

^b Simplified approximate beta-Poisson.

^c Full beta-Poisson.



Fig. 1. Cumulative probability distributions of estimated sewage concentrations under two scenarios: municipal sewage and estimated Davis Station outbreak sewage. The circles are the 50th percentiles and dashed vertical lines are the 95th percentiles.

The required LRVs to meet the $\leq 10^{-6}$ DALYs person⁻¹ year⁻¹ health target, for potable reuse of treated wastewater at Davis Station, were 6.9 for norovirus, 8.0 for giardia and 7.4 for *Campylobacter* using estimates of municipal sewage, while for the Davis Station outbreak scenario they were 12.1, 10.4 and 12.3 respectively (Fig. 2).

The estimate of norovirus concentration in municipal sewage $(1 \times 10^7 \text{ L}^{-1})$ was similar (within 1 order of magnitude) to many previously reported maximum sewage concentrations in Japan, UK. Italy, Finland, Germany, Sweden, Singapore and the Netherlands (Aw and Gin, 2010; Haramoto et al., 2006; Katayama et al., 2008; La Rosa et al., 2010; Laverick et al., 2004; Nordgren et al., 2009; Pusch et al., 2005; Van Den Berg et al., 2005; Von Bonsdorff et al., 2002), while the estimate of outbreak concentration $(1 \times 10^{12} L^{-1})$ was 5 orders of magnitude higher. Similarly, the estimate of giardia concentration in municipal sewage $(9 \times 10^3 \text{ L}^{-1})$ was within 1 order of magnitude of most of the previously reported maximum sewage concentrations in Japan, the Netherlands, Spain, Sweden and the USA (Castro-Hermida et al., 2008; Castro-Hermida et al., 2010; Gassmann and Schwartzbrod, 1991; Medema and Schijven, 2001; Oda et al., 2005; Ottoson et al., 2006a; Ottoson et al., 2006b; Sykora et al., 1991) while the estimate of giardia outbreak concentration $(1.4 \times 10^6 L^{-1})$ was 3 orders of magnitude higher. The estimate of Campylobacter concentration in municipal sewage (7.2×10^3 cfu L⁻¹) was similar to published values from Italy and Spain (Rodríguez and Araujo, 2010; Stellacci et al., 2010), but lower (by as much as 2 orders of magnitude) than published concentrations in Germany and the Baltic Sea region (Holler, 1988; Rechenburg and Kistemann, 2009). The estimate of outbreak concentration (4.9×10^8) was up to 5 orders of magnitude higher than municipal sewage estimates.

The situation considered here is a worst case scenario where raw wastewater is not diluted with other wastewater sources (stormwater, rainwater, etc.). Each of the different scenarios and estimation methods had a significant effect (p < 0.001) on the estimated sewage pathogen concentrations and subsequently the required LRVs. To evaluate the impact of population size on required LRVs, an epidemiological method was developed to estimate norovirus concentrations in Melbourne sewage during an outbreak. Melbourne outbreak sewage concentration ($7.2 \times 10^{10} \text{ H L}^{-1}$) was nearly 4 orders of magnitude greater than municipal sewage ($1.0 \times 10^7 \text{ H L}^{-1}$) and ~1 order of magnitude less than Davis Station outbreak concentration ($1.4 \times 10^{12} \text{ H L}^{-1}$), requiring 10.8 compared with 12.1 LRVs for Davis Station (Fig. 3).

The Guidelines recommend a minimum enteric virus LRV of 9.5 for the production of drinking water from sewage while the model, using municipal sewage pathogen concentrations, determined a LRV of 6.9



Fig. 2. Cumulative probability distributions of required log₁₀ reductions (LRVs) for direct potable reuse of wastewater under two scenarios: municipal sewage and estimated Davis Station outbreak sewage. The circles are the 50th percentiles and dashed vertical lines are the 95th percentiles.



Fig. 3. Cumulative probability distributions of required \log_{10} reductions (LRVs) for direct potable reuse of wastewater comparing estimate methods: 1) municipal sewage, 2) outbreak sewage – Davis Station and 3) outbreak sewage – large city. The circles are the 50th percentiles and dashed vertical lines are the 95th percentiles.

for norovirus. To compare these two methods, sequential steps from the Guideline method to a deterministic approximation of the model are reported in Table 2. The difference in LRVs between steps 2 and 4 shows that the full norovirus dose-response model reduces the required LRV from 12.5 with the rotavirus dose-response model to 7.2; this is likely the primary contributing factor to the difference between the Guideline value and the model value, although the higher virus concentration was also important (increased the LRV from 9.4 to 12.5). The difference between steps 4 and 5 shows the impact of using the higher drinking water volume (7.2 to 7.6) and the difference between steps 5 and 7 shows the impact of a shorter exposure period (7.6 to 7.3); none of these changes greatly altered the final model output. Comparing the 95th percentile of the full stochastic model (6.9) with a deterministic approximation of the model (step 7; 7.3), the difference is small, demonstrating that the understanding gained from the stepwise evaluation of parameter changes can be applied to the full model. A similar step-wise process was conducted for the other reference pathogens and results are presented in supplementary materials (Tables S.4 and S.5). The impact of the full stochastic model had much less impact on the LRVs for giardia and Campylobacter.

An assessment of all input parameters confirmed that there were no unexpected relationships or correlations and variation in many of the input parameters contributed significantly to the variation in the model outputs (Table 3). Using the municipal sewage method, sewage concentration had the largest impact on variation in the estimate of required LRVs, while drinking water volume, disease burden and exposure period contributed smaller amounts. Exposure period did not affect Campylobacter, while for giardia the dose-response parameter (r) and the infection to illness relationship also made significant contributions to variation. The outbreak scenario method was similar for norovirus and Campylobacter, with the greatest effect on variation in LRV due to variation in the estimate of sewage concentration which was a function of the other input parameters. Pathogen shedding rate contributed the most to the variation in LRVs for norovirus and Campylobacter, followed by fecal weight, disease burden, volume of drinking water and daily per capita water use. Secondary attack rate was also a significant contributor for Campylobacter. The variation in required LRVs for giardia was somewhat different and largely influenced by the variation in the dose-response parameter and illness to infection ratio, followed by drinking water volume, exposure period and daily per capita water use.

Similar trends were observed in the impact on LRVs when input parameters were fixed at discrete percentile values (Fig. 4). For municipal sewage scenarios, median LRVs were most affected by the variation in the estimate of sewage concentration, with the spread in estimated LRVs as high as $2.3 \log_{10}$ for giardia. For outbreak sewage scenarios, median LRVs were most affected by pathogen shedding rate for norovirus and *Campylobacter* with a difference in LRVs as large as $1.3 \log_{10}$ (*Campylobacter*). The effect of input parameter variation on LRVs for giardia was minimal for outbreak conditions.

4. Discussion

While there have been recent arguments that the 10^{-6} DALY threshold is too conservative, even for developed countries with lower background levels of water-borne disease (Mara, 2011; Mara et al., 2010), the more cautious approach appears sensible in the context of small communities where, as a result of isolation, the implications of illness may be much greater. Using the 10^{-6} DALY health target, required LRVs were calculated to be 6.9, 8.0 and 7.4 for norovirus, giardia and *Campylobacter* using municipal sewage values and 12.1, 10.4 and 12.3 for estimated Davis Station outbreak conditions, compared with 9.5, 8.0 and 8.1 reported in the Guidelines (NRMMC et al., 2008). Using municipal sewage concentrations, the LRVs for giardia and *Campylobacter* were very similar to the Guideline values while the LRV for norovirus was much lower, largely due to the difference between the rotavirus and norovirus dose–response models.

Under outbreak conditions, LRVs were much higher than Guideline values as a direct result of the much higher sewage pathogen concentrations (3–5 orders of magnitude greater) estimated for Davis Station

Т٦	ы	•	2
Id	IJ	c	э.

Spearman's rank order correlation coefficients for required log₁₀ pathogen reductions.

Pathogen	Method	Model input parameters ^a											
		Csewage	V	$S_{\rm f}$	n	В	r	inf:ill	Р	Ar	S	F	W
Norovirus	Municipal Outbreak	0.90 ^b 0.88 ^{b,c}	0.22 ^b 0.24 ^b	-0.04 - 0.02	0.09 ^b 0.15 ^b	0.28 ^b 0.32 ^b	n/a n/a	n/a n/a	n/a 0.001	n/a 0.09 ^b	n/a 0.75 ^b	n/a 0.41 ^b	n/a -0.21 ^b
Giardia	Municipal Outbreak	0.86 ^b 0.26 ^{b,c}	0.20 ^b 0.34 ^b	n/a ^d n/a	0.16 ^b 0.32 ^b	0.08 ^b 0.12 ^b	0.27 ^b 0.66 ^b	0.23 ^b 0.50 ^b	n/a -0.02	n/a 0.07 ^b	n/a 0.05	n/a n/a	n/a -0.25 ^b
Campylobacter	Municipal Outbreak	0.99 ^b 0.93 ^{b,c}	0.07 ^b 0.23 ^b	n/a n/a	0.00 0.09 ^b	0.14 ^b 0.25 ^b	n/a n/a	n/a n/a	n/a 0.07 ^b	n/a 0.45 ^b	n/a 0.66 ^b	n/a 0.33 ^b	n/a -0.16 ^b

^a Model input parameters: c_{sewage} = estimated sewage pathogen concentration (# L⁻¹), V = daily water consumption (L person⁻¹), S_f = susceptibility fraction, n = exposure period (days year⁻¹), n = exposure period (days year⁻¹), B = disease burden (DALYs case⁻¹), r = dose-response parameter for giardia, *inf:ill* = ratio of infection to illness for giardia, P = station population, A_r = secondary attack rate, S = peak pathogen shedding, F = daily fecal weight (g-feces person⁻¹), W = daily water use (L person⁻¹ day⁻¹). ^b $p \le 0.05$.

^c Outbreak sewage pathogen concentration was calculated from some or all of the following inputs: station population, secondary attack rate, shedding rate, fecal weight, daily water use and dose-response fit parameters. Its inclusion in the sensitivity analysis reflects the sum of variation contributed by the other model input parameters. ^d n/a = not applicable.



Fig. 4. Median required \log_{10} reductions when individual input parameters were held at discrete values. Boxes represent values for 5th, 50th and 95th percentile input values (bottom, middle and upper lines, respectively). The left hand side depicts municipal sewage scenarios and the right hand side depicts the Davis Station outbreak scenarios. Input parameters are defined as follows: c = sewage concentration, V = daily water consumption, n = exposure period, B = disease burden (DALYs case⁻¹), Sf = susceptibility fraction, r = dose-response parameter, *inf:ill* = ratio of infection to illness, P = station population, W = daily water, Ar = secondary attack rate, *inf* = number of people ill, *SR* = shedding rate, f = daily fecal weight.

outbreak conditions. These values, particularly norovirus, were orders of magnitude higher than other published values of municipal sewage pathogen concentrations, reporting peaks of 10³-10⁷ for norovirus, $10^2 - 10^4$ for giardia and $10^4 - 10^7$ for *Campylobacter* (Table S.6). There is very little information available on sewage pathogen concentrations during community gastroenteritis outbreaks, although the Guidelines use 95th percentile values assumedly to represent peak pathogen loads that might occur during an outbreak. To further evaluate the outbreak method, norovirus concentrations at Davis Station were compared with estimated concentrations during an outbreak in Melbourne. The proportion of people that become infected during a Melbourne norovirus outbreak (1%) was much less than the secondary attack rate (14-18%) used for the Davis Station outbreak scenario; therefore, Melbourne sewage was more dilute (i.e. lower pathogen concentration) and required 10.8 compared with 12.1 LRVs for Davis. Assuming that the 95th percentile of the municipal concentration estimate represents outbreak conditions, the median Melbourne outbreak concentration $(2.6 \times 10^{10} \text{ # L}^{-1})$ was nearly 3 orders of magnitude higher and may represent an overestimation of outbreak concentrations. There are various possible explanations for this disparity in concentration estimates: 1) the estimate of municipal sewage, based on data from Japan, does not reflect Melbourne conditions (i.e. norovirus rates in Japan are lower than in Melbourne); 2) the estimate of municipal sewage, based on monthly measurements, missed outbreak conditions; 3) the outbreak method does not account for pathogen decay through the distribution system; or 4) the outbreak sewage estimation method is too conservative. The impact of each of these potential contributors cannot be quantified but importantly, even if the outbreak method overestimates sewage concentration, the required LRVs are still higher than those in the Guidelines suggesting that additional treatment will be required. A greater understanding of sewage pathogen concentrations from small communities is needed to reduce the uncertainty around the estimated LRVs.

Various assumptions were made in the development of the model that may be important constraints in the application of the model results. Secondary attack rate was used to estimate outbreak sewage pathogen concentrations and is a measure of the spread of illness by direct (person-to-person contact, inhalation of aerosols, etc.) and indirect (transfer from contaminated surfaces, etc.) contact. Studies are typically conducted in relatively confined populations such as households and school camps. While there is evidence that pathogen shedding can occur in the absence of symptoms (Atmar et al., 2008; Birkhead and Vogt, 1989; Yakoob et al., 2010), the secondary attack rate accounts for symptomatic cases only. Therefore, the model has not accounted for asymptomatic infections that could contribute to the pathogen load in sewage. This may be of limited concern, at least for norovirus, as recent investigations have found that asymptomatic cases are unlikely to cause transmission despite high shedding rates (Sukhrie, 2012). We have also made highly conservative assumptions that all individuals became ill instantaneously and shed pathogens at the peak rate, and that all infected or ill individuals had diarrhea. In an actual outbreak, it is likely that the spread of infection would occur over a few weeks (the time span of studies used to estimate secondary attack rate). At the same time, pathogen shedding can occur for extended periods of time – both prior to symptomatic illness and after apparent recovery - and it would seem unlikely that peak shedding amongst all individuals would occur simultaneously.

Careful consideration will be required to design a treatment plant to meet safe drinking water requirements in the event of an outbreak of gastroenteritis in a small community. The higher required LRVs for norovirus, giardia and Campylobacter will demand a combination of treatment systems. At Davis Station, a secondary treatment plant will be installed to remove the majority of the wastewater contaminants, with additional tertiary and polishing treatment steps to meet potable water quality requirements. The tertiary and polishing processes of large scale indirect potable water systems generally consist of ultrafiltration, reverse osmosis and advanced oxidation followed by final disinfection. Such systems provide a multi-barrier approach to ensure water quality and are required to achieve a virus LRV of 9.5. Such processes can achieve higher LRVs (e.g. virus LRV of 10 for Western Corridor in Brisbane, Australia), but nevertheless, the higher required LRVs for small scale treatment plants as suggested by this model (e.g. an extra LRV of 2.6 for viruses) will necessitate additional treatment units such as UV disinfection. The higher protozoa and bacteria LRVs required for small systems also necessitate this extra treatment barrier.

In considering the higher required LRV requirements suggested by this model, it is important to contextualize the risk of exposure to treated wastewater relative to other forms of exposure. A small community such as Davis Station operates similar to a household in that the level of contact between community members is quite high. The potential exposure pathways include person-to-person contact, contact with contaminated surfaces and inhalation/ingestion of aerosols. The assumption of the model, that one infected person arrives at Davis Station, would result in 18, 19 or 12 people sick with norovirus, giardia or *Campylobacter* respectively, based on the secondary attack rate (direct or indirect contact with the infected person). In contrast, assuming all infected individuals are shedding pathogens at a peak rate and that treatment of sewage conforms to the required LRVs needed to meet the 10^{-6} DALY health target, consumption of the treated water would result in up to 17 cases of norovirus, 5 cases of giardia or 2 cases of Campylobacter illness per 10,000 people or 0.18, 0.05 and 0.02 additional cases of norovirus, giardia and Campylobacter per summer season (using 95th percentile station population).

While Davis Station may be considered an extreme example, a similar approach could be applied to many small remote communities in Australia. In the Northern Territory alone, there are 41 predominantly indigenous communities (95% indigenous) that range in size from 85 to 886 residents, with 13 of those communities having a population under 200 (ABS, 2007b). Other reports have found that of the 1,139 remote indigenous communities across Australia, more than half (54%) reported less than 20 residents and 23% reported populations of 20 to 49 (ABS, 2003). DPR may be an appropriate solution in some of these communities and the results of this model demonstrate the importance of consideration of small communities in determining appropriate treatment trains.

5. Conclusion

Direct potable reuse is a relatively new concept that has legitimate potential to enhance water security in both small and large communities. This analysis has highlighted the need to consider population size and vulnerability when assessing treatment requirements, a conclusion based on a quantitative microbial risk assessment (QMRA) that was conducted using norovirus, giardia and Campylobacter as reference pathogens. Two scenarios were compared, municipal sewage pathogen loads and potential pathogen loads during a community gastroenteritis outbreak, and pathogen concentrations were significantly higher (p < 0.001) in the outbreak scenario. For the municipal sewage scenario, required LRVs were 6.9, 8.0 and 7.4 for norovirus, giardia and *Campylobacter* respectively, while for outbreak conditions, the values were 12.1, 10.4 and 12.3. While the outbreak values could overestimate LRVs by as much as 3 (for norovirus), they still indicate a need for additional treatment barriers for small communities in order to provide safe drinking water in the event of an outbreak. This higher treatment requirement is predominately attributed to the significantly increased pathogen levels in outbreak sewage relative to municipal sewage from a large city as a result of dilution and the relatively smaller proportion of the population infected. The recommended pathogen LRVs clearly represent a worst case scenario, assuming high pathogen concentrations and close community contact (high secondary attack rate). Generalization to other small communities is relevant nonetheless, and the model results indicate that in the event of an outbreak additional treatment barriers will be necessary to achieve safe drinking water in such communities.

Conflict of interest

There is no conflict of interest to report.

Acknowledgments

The authors would like to acknowledge the members of the AAD Polar Medicine Unit for providing context and a clearer understanding of station conditions. Helpful comments and suggestions were also received from Dr. Martha Sinclair (Monash University). Funding for this work through the Cooperative Research Network (CRN) program of the Australian Government and the Australian Antarctic Division of the Department of Sustainability, Environment, Water, Populations and Communities of the Australian Government is gratefully acknowledged.

Appendix A. Supplementary data

Supplementary data to this article can be found online at http://dx.doi.org/10.1016/j.scitotenv.2013.05.059.

References

- AAD. Antarctic explorer expeditioner's database. Kingston, Tasmania: Australian Antarctic Division; 2011.
- AAD. State of environment. Indicator 61 total potable water consumption at Australian Antarctic Stations. SIMR — (state of environment). System for indicator management and reporting — an on-line state of environment system for the Antarctic. SIMR — (state of environment). Australian Antarctic Division (AAD): 2012.

ABS. 4102.0 Australian social trends 2003. Canberra, Australia: Australian Bureau of Statistics; 2003.

- ABS. 4602.0 Environmental issues: people's views and practices, Mar 2007. Commonwealth of Australia, Canberra: Australian Bureau of Statistics; 2007a.
- ABS. 4705.0 Population distribution, aboriginal and Torres Strait Islander Australians, 2006. Canberra, Australia: Australian Bureau of Statistics; 2007b.
- ABS. 4102.0 Australian social trends 2008. Commonwealth of Australia, Canberra: Australian Bureau of Statistics; 2008.
- ABS. 4602.0.55.003 environmental issues: water use and conservation. Commonwealth of Australia, Canberra: Australian Bureau of Statistics; 2010a.
- ABS. 4602.2 household water, energy use and conservation, Victoria, Oct 2009. Commonwealth of Australia, Canberra: Australian Bureau of Statistics; 2010b.
- ABS. 3101.0 Australian demographic statistics, Mar 2012. Canberra: Australian Bureau of Statistics; 2012.
- Alfano-Sobsey E, Sweat D, Hall A, Breedlove F, Rodriguez R, Greene S, et al. Norovirus outbreak associated with undercooked oysters and secondary household transmission. Epidemiol Infect 2012;140:276–82.
- Astrom J, Petterson S, Bergstedt O, Pettersson TJR, Stenström TA. Evaluation of the microbial risk reduction due to selective closure of the raw water intake before drinking water treatment. J Water Health 2007;5:81–97.
- Atmar RL. Noroviruses: state of the art. Food Environ Virol 2010;2:117-26.
- Atmar RL, Opekun AR, Gilger MA, Estes MK, Crawford SE, Neill FH, et al. Norwalk virus shedding after experimental human infection. Emerg Infect Dis 2008;14:1553–7. Aw TG, Gin KYH. Environmental surveillance and molecular characterization of human
- enteric viruses in tropical urban wastewaters. J Appl Microbiol 2010;109:716–30. Baron RC, Murphy FD, Greenberg HB, Davis CE, Bregman DJ, Gary GW, et al. Norwalk gastrointestinal illness: an outbreak associated with swimmin in a recreational lake and secondary person-to-person transmission. Am J Epidemiol 1982;115:
- 163–72.Birkhead G, Vogt RL. Epidemiologic surveillance for endemic *Giardia lamblia* infection in vermont the roles of waterborne and person-to-person transmission.
- Am J Epidemiol 1989;129:762–8. Buckland ST. Monte carlo confidence intervals. Biometrics 1984;40:811–7.
- Castro-Hermida JA, García-Presedo I, Almeida A, González-Warleta M, Da Costa JMC, Mezo M. Contribution of treated wastewater to the contamination of recreational river areas with *Cryptosporidium* spp. and *Giardia duodenalis*. Water Res 2008;42: 3528–38.
- Castro-Hermida JA, García-Presedo I, González-Warleta M, Mezo M. Cryptosporidium and Giardia detection in water bodies of Galicia, Spain. Water Res 2010;44: 5887–96.
- Chan MCW, Sung JJY, Lam RKY, Chan PKS, Lee NLS, Lai RWM, et al. Fecal viral load and norovirus-associated gastroenteritis. Emerg Infect Dis 2006;12:1278–80.
- Cressey P, Lake R. Risk ranking: DALY estimates for selected foodborne diseases in New Zealand using revised Dutch disability weights. Ilam, Christchurch: Institute of Environmental Science & Research Limited; 2009.
- Denborough MA, Downing HJ. Secretor status in asthma and hay fever. J Med Genet 1968;5:302–5.
- Evans MR. A milk-borne campylobacter outbreak following an educational farm visit. Epidemiol Infect 1996;117:457–62.
- Feachem RG, Bradley DJ, Garelick H, Mara DD. Sanitation and disease. Health aspects of excreta and wastewater management. Washington, D.C.: John Wiley & Sons; 1983
- Gassmann L, Schwartzbrod J. Wastewater and giardia cysts. Water Sci Technol 1991;24: 183–6.
- Götz H, De Jong B, Lindbäck J, Parment PA, Hedlund KO, Torvén M, et al. Epidemiological investigation of a food-borne gastroenteritis outbreak caused by Norwalk-like virus in 30 day-care centres. Scand J Infect Dis 2002;34:115–21.
- Haagsma JA, Havelaar AH, Janssen BMF, Bonsel GJ. Disability adjusted life years and minimal disease: application of a preference-based relevance criterion to rank enteric pathogens. Popul Health Metr 2008:6.
- Hall GV, Kirk MD, Ashbolt R, Stafford R, Lalor K, Bell R, et al. Frequency of infectious gastrointestinal illness in Australia, 2002: regional, seasonal and demographic variation. Epidemiol Infect 2006;134:111–8.
- Haramoto E, Katayama H, Oguma K, Yamashita H, Tajima A, Nakajima H, et al. Seasonal profiles of human noroviruses and indicator bacteria in a wastewater treatment plant in Tokyo, Japan. Water Sci Technol 2006;54:301–8.
- Havelaar AH. Disease burden of foodborne pathogens in the Netherlands, 2009. Int J Food Microbiol 2012;156:231–8.
- Havelaar AH, Melse JM. Quantifying public health risk in the WHO Guidelines for drinking water quality. A burden of disease approach. Bilthoven, BA: World Health Organization, Netherlands Ministry of Housing, Physical Planning and the Environment, Directorate General for Environmental Protection, Directorate for Soil, Water and Countryside; 2003.
- Holler C. Long-term study of occurrence, distribution and reduction of *Campylobacter* sp. in the sewage system and wastewater treatment plant of a big town. Water Sci Technol 1988;20:529–31.
- Hoque EM, Hope VT, Kjellström T, Scragg R, Lay-Yee R. Risk of giardiasis in Aucklanders: a case—control study. Int J Infect Dis 2002;6:191–7.
- Hunter PR, De Sylor MA, Risebro HL, Nichols GL, Kay D, Hartemann P. Quantitative microbial risk assessment of *Cryptosporidiosis* and *Giardiasis* from very small private water supplies. Risk Anal 2011;31:228–36.
- Johansson PJH, Torvén M, Hammarlund AC, Björne U, Hedlund KO, Svensson L. Food-borne outbreak of gastroenteritis associated with genogroup I calicivirus. J Clin Microbiol 2002;40:794–8.
- Johnson PC, Mathewson JJ, DuPont HL, Greenberg HB. Multiple-challenge study of host susceptibility to Norwalk gastroenteritis in US adults. J Infect Dis 1990;161:18–21.

- Katayama H, Haramoto E, Oguma K, Yamashita H, Tajima A, Nakajima H, et al. One-year monthly quantitative survey of noroviruses, enteroviruses, and adenoviruses in wastewater collected from six plants in Japan. Water Res 2008;42:1441–8.
- Katz DE, Heisey-Grove D, Beach M, Dicker RC, Matyas BT. Prolonged outbreak of giardiasis with two modes of transmission. Epidemiol Infect 2006;134:935–41.
- Kemmeren JM, Mangen M-JJ, van Duynhoven YTHP, Havelaar AH. Priority setting of foodborne pathogens: disease burden and costs of selected enteric pathogens (RIVM report 330080001/2006). Bilthoven, The Netherlands: National Institute for Public Health and the Environment (RIVM); 2006.
- La Rosa G, Pourshaban M, Iaconelli M, Muscillo M. Quantitative real-time PCR of enteric viruses in influent and effluent samples from wastewater treatment plants in Italy. Ann Ist Super Sanita 2010;46:266–73.
- Lahnsteiner J, Lempert G. Water management in Windhoek, Namibia. Water Sci Technol 2007;55:441–8.
- Lake RJ, Cressey PJ, Campbell DM, Oakley E. Risk ranking for foodborne microbial hazards in New Zealand: burden of disease estimates. Risk Anal 2010;30:743–52.
- Laverick MA, Wyn-Jones AP, Carter MJ. Quantitative RT-PCR for the enumeration of noroviruses (Norwalk-like viruses) in water and sewage. Lett Appl Microbiol 2004;39: 127–36.
- Le Pendu J, Ruvoen-Clouet N, Kindberg E, Svensson L. Mendelian resistance to human norovirus infections. Semin Immunol 2006;18:375–86.
- Lee N, Chan MCW, Wong B, Choi KW, Sin W, Lui G, et al. Fecal viral concentration and diarrhea in norovirus gastroenteritis. Emerg Infect Dis 2007;13:1399–401.
- Leverenz HL, Tchobanoglous G, Asano T. Direct potable reuse: a future imperative. J Water Reuse Desalination 2011;1:2–10.
- Lin S, Wang X, Zheng H, Mao Z, Sun Y, Jiang B. Direct detection of *Campylobacter jejuni* in human stool samples by real-time PCR. Can J Microbiol 2008;54:742–7.
- Lindesmith L, Moe C, Marionneau S, Ruvoen N, Jiang X, Lindblad L, et al. Human susceptibility and resistance to Norwalk virus infection. Nat Med 2003;9:548–53.
- Lopez CE, Dykes AC, Juranek DD. Waterborne giardiasis: a communitywide outbreak of disease and a high rate of asymptomatic infection. Am J Epidemiol 1980;112:495–507.
- Mara DD. Water- and wastewater-related disease and infection risks: what is an appropriate value for the maximum tolerable additional burden of disease? J Water Health 2011;9:217–24.
- Mara DD, Hamilton AJ, Sleigh A, Karavarsamis N. Discussion paper: options for updating the 2006 WHO guidelines. WHO, FAO, IDRC, IWMI; 2010.
- Masago Y, Katayama H, Watanabe T, Haramoto E, Hashimoto A, Omura T, et al. Quantitative risk assessment of noroviruses in drinking water based on qualitative data in Japan. Environ Sci Technol 2006;40:7428–33.
- Medema GJ, Schijven JF. Modelling the sewage discharge and dispersion of Cryptosporidium and giardia in surface water. Water Res 2001;35:4307–16.
- Nordgren J, Matussek A, Mattsson A, Svensson L, Lindgren PE. Prevalence of norovirus and factors influencing virus concentrations during one year in a full-scale wastewater treatment plant. Water Res 2009;43:1117–25.
- Norkrans G, Svedhem Å. Epidemiological aspects of Campylobacter jejuni enteritis. J Hyg 1982;89:163–70.
- NRMMC, EPHC, AHMC. National guidelines for water recycling: managing health and environmental risks (phase 1). National water quality management strategy. Natural Resource Management Ministerial Council, Environment Protection and Heritage Council. Canberra: Australian Health Ministers' Conference; 2006a.
- NRMMC, EPHC, AHMC. Australian guidelines for water recycling: managing health and environmental risks (phase 1). National water quality management strategy. Natural Resource Management Ministerial Council, Environment Protection and Heritage Council, Australian Health Ministers' Conference, Canberra; 2006b.
- NRMMC, EPHC, NHMRC. Australian guidelines for water recycling: managing health and environmental risks (Phase 2). Augmentation of drinking water supplies. National Water Quality Management Strategy. Canberra: Natural Resource Management Ministerial Council, Environment Protection and Heritage Council, National Health and Medical Research Council; 2008.
- NRMMC, EPHC, NHMRC. Australian guidelines for water recycling: managing health and environmental risks (phase 2). Managed aquifer recharge. National water quality management strategy. Canberra: Natural Resource Management Ministerial Council, Environment Protection and Heritage Council, National Health and Medical Research Council; 2009.
- Oda T, Kawabata M, Uga S. Detection of giardia cysts in sewage and estimations of giardiasis prevalence among inhabitants in Hyogo Prefecture, Japan. Trop Med Health 2005;33:1–5.
- Ottoson J, Hansen A, Bjorlenius B, Norder H, Stenström TA. Removal of viruses, parasitic protozoa and microbial indicators in conventional and membrane processes in a wastewater pilot plant. Water Res 2006a;40:1449–57.
- Ottoson J, Hansen A, Westrell T, Johansen K, Norder H, Stenström TA. Removal of noro- and enteroviruses, *Ciardia cysts*, *Cryptosporidium* oocysts, and fecal indicators at four secondary wastewater treatment plants in Sweden. Water Environ Res 2006b;78:828–34.
- Pickering LK, Evans DG, DuPont HL. Diarrhea caused by Shigella, rotavirus, and Giardia in day-care centers: prospective study. J Pediatr 1981;99:51–6.
- Pintar KDM, Fazil A, Pollari F, Waltner-Toews D, Charron DF, McEwen SA, et al. Considering the risk of infection by *Cryptosporidium* via consumption of municipally treated drinking water from a surface water source in a southwestern Ontario community. Risk Anal 2012;32:1122–38.
- Porter IA, Reid TMS. A milk-borne outbreak of *Campylobacter* infection. J Hyg 1980;84: 415–9.
- Pusch D, Oh DY, Wolf S, Dumke R, Schröter-Bobsin U, Höhne M, et al. Detection of enteric viruses and bacterial indicators in German environmental waters. Arch Virol 2005;150:929–47.
- Rao SSC. Oral rehydration for viral gastroenteritis in adults: a randomized, controlled trial of 3 solutions. JPEN J Parenter Enteral Nutr 2006;30:433–9.

- Rechenburg A, Kistemann T. Sewage effluent as a source of *Campylobacter* sp. in a surface water catchment. Int J Environ Health Res 2009;19:239–49.
- Roche SM, Jones AQ, Majowicz SE, McEwen SA, Pintar KDM. Drinking water consumption patterns in Canadian communities (2001–2007). J Water Health 2012;10:69–86.
- Rodríguez S, Araujo R. Occurrence of thermotolerant Campylobacter species in surface waters of a Mediterranean area and in its prevailing pollution sources. J Appl Microbiol 2010;109:1027–34.
- Rosström-Lindquist K, Palm D, Reiner D, Ringqvist E, Svärd SG. Giardia immunity an update. Trends Parasitol 2006;22:26–31.
- Schijven JF, Teunis PFM, Rutjes SA, Bouwknegt M, de Roda Husman AM. QMRAspot: a tool for quantitative microbial risk assessment from surface water to potable water. Water Res 2011;45:5564–76.
- Sinclair MI, Hellard ME, Wolfe R, Mitakakis TZ, Leder K, Fairley CK. Pathogens causing community gastroenteritis in Australia. J Gastroenterol Hepatol 2005;20:1685–90.
- Soller JA, Bartrand T, Ashbolt NJ, Ravenscroft J, Wade TJ. Estimating the primary etiologic agents in recreational freshwaters impacted by human sources of faecal contamination. Water Res 2010;44:4736–47.
- Stellacci P, Liberti L, Notarnicola M, Haas CN. Hygienic sustainability of site location of wastewater treatment plants. A case study. II. Estimating airborne biological hazard. Desalination 2010;253:106–11.
- Sukhrie FHA. Nosocomial transmission of norovirus is mainly caused by symptomatic cases. Clin Infect Dis 2012;54:931–7.
- Sykora JL, Sorber CA, Jakubowski W, Casson LW, Gavaghan PD, Shapiro MA, et al. Distribution of *Giardia* cysts in wastewater. Water Sci Technol 1991;24:187–92.
- Tchobanoglous G, Leverenz H, Nellor MH, Crook J. Direct potable reuse: a path forward. Alexandria, VA: WateReuse Research Foundation; 2011.
- ter Waarbeek HLG, Dukers-Muijrers NHTM, Vennema H, Hoebe CJPA. Waterborne gastroenteritis outbreak at a scouting camp caused by two norovirus genogroups: GI and GII. J Clin Virol 2010;47:268–72.
- Teunis P, van der Heijden O, van der Giessen J, Havelaar A. The dose-response relation in human volunteers for gastro-intestinal pathogens. Report nr. 284550002. Bilthoven, The Netherlands: National Institute of Public Health and the Environment (RIVM); 1996.
- Teunis PFM, van den Brandhof W, Nauta M, Wagenaar J, van den Kerkhof H, van Pelt W. A reconsideration of the *Campylobacter* dose–response relation. Epidemiol Infect 2005;133:583–92.

Teunis PFM, Moe CL, Liu P, Miller SE, Lindesmith L, Baric RS, et al. Norwalk virus: how infectious is it? J Med Virol 2008;80:1468–76.

- The R Foundation for Statistical Computing. The R project for statistical computing; 2011. Thorven M, Grahn A, Hedlund KO, Johansson H, Wahlfrid C, Larson G, et al. A homozygous nonsense mutation (428G-A) in the human secretor (FUT2) gene provides resistance to symptomatic norovirus (GGII) infections. J Virol 2005;79:15351–5.
- Tsuchiya H. A study on variabilities in dimensions and numbers of discharged cysts *Giardia lamblia* (stiles 1915) from day to day under normal conditions. Am J Hyg 1931;13:544–67.
- USEPA. Estimated per capita water ingestion and body weight in the United States an update. Washington, D C: U.S. EPA, Office of Water, Office of Science and Technology; 2004.
- USEPA. Economic analysis for the final ground water rule. United States Environmental Protection Agency; 2006.
- Van Den Akker B, Whiffin V, Cox P, Beatson P, Ashbolt NJ, Roser DJ. Estimating the risk from sewage treatment plant effluent in the Sydney catchment area. Water Sci Technol 2011;63:1707–15.
- Van Den Berg H, Lodder W, Van Der Poel W, Vennema H, De Roda Husman AM. Genetic diversity of noroviruses in raw and treated sewage water. Res Microbiol 2005;156: 532–40.
- Victoria M, Guimarģes FR, Fumian TM, Ferreira FFM, Vieira CB, Leite JPG, et al. One year monitoring of norovirus in a sewage treatment plant in Rio de Janeiro, Brazil. J Water Health 2010;8:158–65.
- Vijgen SMC, Mangen MJM, Kortbeek LM, van Duijnhoven YTHP, Havelaar AH. Disease burden and related costs of cryptosporidium and giardiasis in the Netherlands. Bilthoven: RIVM; 2007.
- Von Bonsdorff CH, Maunula L, Niemi RM, Rimhanen-Finne R, Hänninen ML, Lahti K. Hygienic risk assessment by monitoring pathogens in municipal sewage. Water Sci Technol 2002;2:23–8.
- WHO. Guidelines for the safe use of wastewater, excreta and greywater. Geneva, Switzerland: World Health Organisation; 2006.
- WHO. Guidelines for drinking water quality. 4th ed Geneva: World Health Organization; 2011.
- Yakoob J, Abbas Z, Beg MA, Naz S, Khan R, Islam M, et al. Prevalences of *Giardia lamblia* and *Cryptosporidium parvum* infection in adults presenting with chronic diarrhoea. Ann Trop Med Parasitol 2010;104:505–10.